

Aufgaben zu Kapitel 4

26.5.26

Die Zufallsvariable

$$t = \frac{\hat{\beta}_j - \beta_j}{\sqrt{\widehat{\text{Var}}(\hat{\beta}_j | X)}}$$

ist t -verteilt.
mit $n - k - 1$ Freiheitsgraden

Im Allgemeinen:

$$X \sim N(0, 1)$$

$$Y \sim \chi^2_{n-k-1}$$

X, Y unabhängig

$$\rightarrow \frac{X}{\sqrt{Y/n-k-1}}$$

t -verteilt mit $n - k - 1$ Freiheitsgraden

$$t = \frac{\hat{\beta}_j - \beta_j}{\sqrt{\widehat{\text{Var}}(\hat{\beta}_j | X)}}$$

$$\frac{\hat{\beta}_j - \beta_j}{\sqrt{\text{Var}(\hat{\beta}_j | X)}}$$

$\sim \mathcal{N}(0, 1)$ Aufgabe 1

$$\frac{\sqrt{\widehat{\text{Var}}(\hat{\beta}_j | X)}}{\sqrt{\text{Var}(\hat{\beta}_j | X)}}$$

= " $\sqrt{\frac{y}{n-k-1}}$ " Aufgabe 2

1 $V \sim \mathcal{N}(\mu, \sigma_v^2)$

$$\frac{\hat{\beta}_j - \beta_j}{\sqrt{\text{Var}(\hat{\beta}_j)}} \sim \mathcal{N}(0, 1)$$

Zufällig \downarrow $V - \mu$ \leftarrow fest

$$z = \frac{V - \mu}{\sqrt{\sigma_v^2}} \sim \mathcal{N}(E[z], \text{Var}(z))$$

\leftarrow fest

$$\leadsto z \sim \mathcal{N}(0, 1)$$

$$E[z] = E\left[\frac{V - \mu}{\sqrt{\sigma_v^2}}\right] = \frac{1}{\sqrt{\sigma_v^2}} E[V - \mu] = \frac{1}{\sqrt{\sigma_v^2}} (\underbrace{E[V]}_{=\mu} - \mu) = 0$$

$$\text{Var}(z) = \text{Var}\left(\frac{V - \mu}{\sqrt{\sigma_v^2}}\right) = \left(\frac{1}{\sqrt{\sigma_v^2}}\right)^2 \cdot \text{Var}(V - \mu) = \frac{1}{\sigma_v^2} \underbrace{\text{Var}(V)} = \frac{1}{\sigma_v^2} \cdot \sigma_v^2 = 1$$

$$\frac{1}{\sigma} u \sim \mathcal{N}(0, I_n)$$

$\uparrow \qquad \nwarrow$
 $E\left[\frac{u}{\sigma}\right] \quad \mathcal{F}\left(\frac{u}{\sigma}\right)$

MLR 6: $u \sim \mathcal{N}(0, \sigma^2 I)$

$$E\left[\frac{1}{\sigma} \cdot u\right] = \frac{1}{\sigma} \cdot E[u] = \frac{1}{\sigma} \cdot 0 = 0 \quad \checkmark$$

$$\mathcal{F}\left(\frac{1}{\sigma} \cdot u\right) = \left(\frac{1}{\sigma}\right)^2 \cdot \mathcal{F}(u) = \frac{1}{\sigma^2} \cdot \sigma^2 \cdot I = I \quad \checkmark$$

$$\# 2 \quad M_x = I - X(X'X)^{-1}X'$$

$$\begin{aligned} \hat{U} &= y - \hat{y} = y - X\hat{\beta} = y - X(X'X)^{-1}X'y \\ &= \underbrace{\left[I - X(X'X)^{-1}X' \right]}_{M_x} y \end{aligned}$$

$$\hat{U} = M_x \cdot y = \dots = M_x \cdot U$$

$$y = X\beta + U$$

M_x idempotent
symmetrisch

$$\hat{U}'\hat{U} = (M_x U)' M_x U = U' \underbrace{M_x' M_x}_{M_x} U = U' M_x U$$

$$\begin{aligned} (n-k-1) \frac{\hat{\sigma}^2}{\sigma^2} &= U' M_x U \frac{1}{\sigma^2} \\ &= \left(\frac{U}{\sigma} \right)' M_x \frac{U}{\sigma} \end{aligned}$$

$$\hat{\sigma}^2 = \frac{1}{n-k-1} \cdot \underbrace{\hat{U}'\hat{U}}_{SSR}$$

$$\left(\frac{U}{\sigma}\right)' \left(\frac{U}{\sigma}\right) = \sum_{i=1}^n \left(\frac{U_i}{\sigma}\right)^2 \sim \chi_n^2$$

$$\left(\frac{U}{\sigma}\right)' M_x \frac{U}{\sigma} \sim \chi_{n-k-1}^2$$

↑
idempotent symmetrisch

$$\text{rk}(M_x) = n - k - 1 = \text{tr}(M_x)$$

$$\frac{1}{n-k-1} \cdot \vec{v}' \vec{v}$$

Also $\left(\frac{U}{\sigma}\right)' \frac{U}{\sigma} = \frac{1}{\sigma^2} \cdot \vec{v}' \vec{v} = \frac{1}{\sigma^2} \cdot (n-k-1) \cdot \boxed{\frac{1}{\sigma^2}}$

$$(n-k-1) \frac{1}{\sigma^2} \sim \chi_{n-k-1}^2$$

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$$\frac{\sqrt{\text{Var}(\hat{\beta}_j | x)}}{\sqrt{\text{Var}(\beta_j | x)}} = \frac{\sqrt{\hat{\sigma}^2 (x'x)^{-1}_{j,j}}}{\sqrt{\sigma^2 (x'x)^{-1}_{j,j}}} = \frac{\sqrt{\hat{\sigma}^2}}{\sqrt{\sigma^2}}$$

$$= \sqrt{\frac{\hat{\sigma}^2}{\sigma^2}} = \sqrt{\frac{(n-k-1) \frac{\hat{\sigma}^2}{\sigma^2}}{(n-k-1)}} \sim \chi^2_{n-k-1}$$

$$t_j = \frac{\hat{\beta}_j - \beta_j}{\sqrt{\widehat{\text{Var}}(\hat{\beta}_j | x)}} = \frac{\overset{x}{\frac{\hat{\beta}_j - \beta_j}{\sqrt{\text{Var}(\hat{\beta}_j | x)}}}}{\sqrt{\frac{\underset{y}{(n-k-1) \hat{\sigma}^2}}{\sigma^2} / (n-k-1)}} \sim N(0,1)$$

$\sim \chi^2_{n-k-1}$

Als letztes wäre zu zeigen: x & y unabhängig

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Modell 1: KQ, benutze die Beobachtungen 1-209

Abhängige Variable: lsalary

	Koeffizient	Std.-fehler	t-Quotient	p-Wert
$\hat{\beta}_0$ const	4,31171	0,315433	13,67	1,16e-30 ***
$\hat{\beta}_1$ lsales	0,280315	0,0353200	7,936	1,34e-13 ***
$\hat{\beta}_2$ roe	0,0174168	0,00409230	4,256	3,17e-05 ***
$\hat{\beta}_3$ ros	<u>0,000241655</u>	<u>0,000541802</u>	<u>0,4460</u>	<u>0,6561</u>
Mittel abhängige Var.	6,950386	Stdabw. abhängige Var.	0,566374	
Summe quad. Residuen	47,86083	Stdfehler Regression	0,483185	
R-Quadrat	0,282685	Korrigiertes R-Quadrat	0,272188	
F(3, 205)	26,92930	P-Wert(F)	1,00e-14	
Log-Likelihood	-142,5213	Akaike-Kriterium	293,0426	
Schwarz-Kriterium	306,4119	Hannan-Quinn-Kriterium	298,4479	

Abgesehen von Konstante war p-Wert am höchsten für Variable 6 (ros)

Test-Entsch.

$$|t_3| < C$$

$\Rightarrow H_0: \beta_3 = 0$
nicht ablehnen

ros \nearrow 50

\rightarrow salary

$$\nearrow \hat{\beta}_3 \%$$

$$0,00024 \cdot 50$$

$$\approx 0,012 \%$$

$$H_0: \beta_3 = 0 \Rightarrow \hat{\beta}_3 \sim N(0, \text{Var}(\hat{\beta}_3 | x))$$

Fehlerw'keit : α

$$t_3 = \frac{\hat{\beta}_3 - 0}{\sqrt{\widehat{\text{Var}}(\hat{\beta}_3)}} \sim t_{205}$$

se($\hat{\beta}_3$)

suche C sodass

$$C_{t_{120}; 99,5\%} = 2,617$$

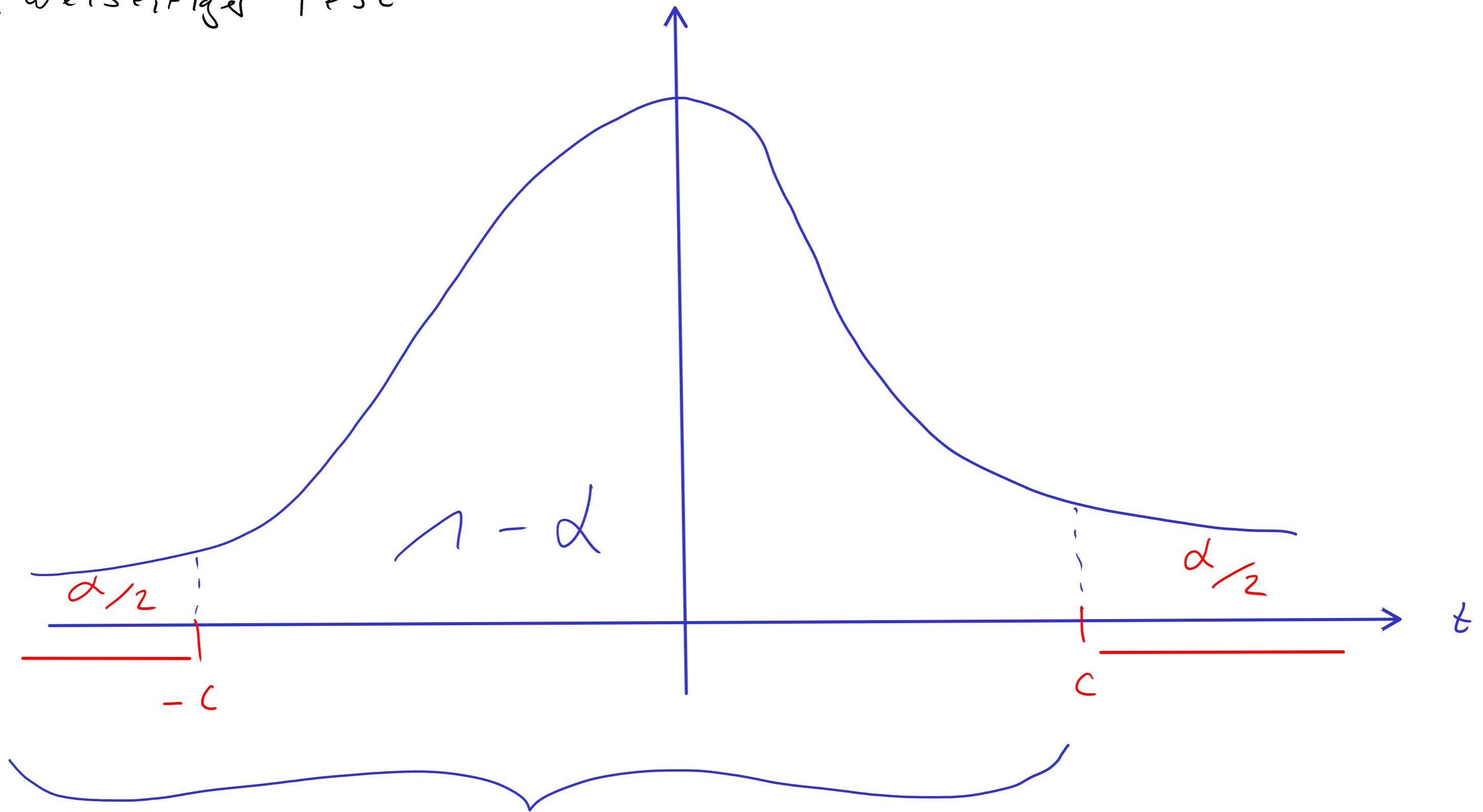
$$\text{Prob}(t_3 \geq C) = \frac{\alpha}{2}$$

$$= \underline{0,4460}$$

$$\Leftrightarrow \text{Prob}(t_3 < -C) = \frac{\alpha}{2} \Leftrightarrow \text{Prob}(-C \leq t_3 \leq C) = 1 - \alpha$$

$$1,16 e^{-30} = 1,16 \cdot 10^{-30}$$

Zweiseitiger Test



$$1 - \alpha + \frac{\alpha}{2} = 1 - \frac{\alpha}{2}$$

$$\text{Prob}(t \leq c) = 1 - \frac{\alpha}{2}$$

noch

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Einseitiger test

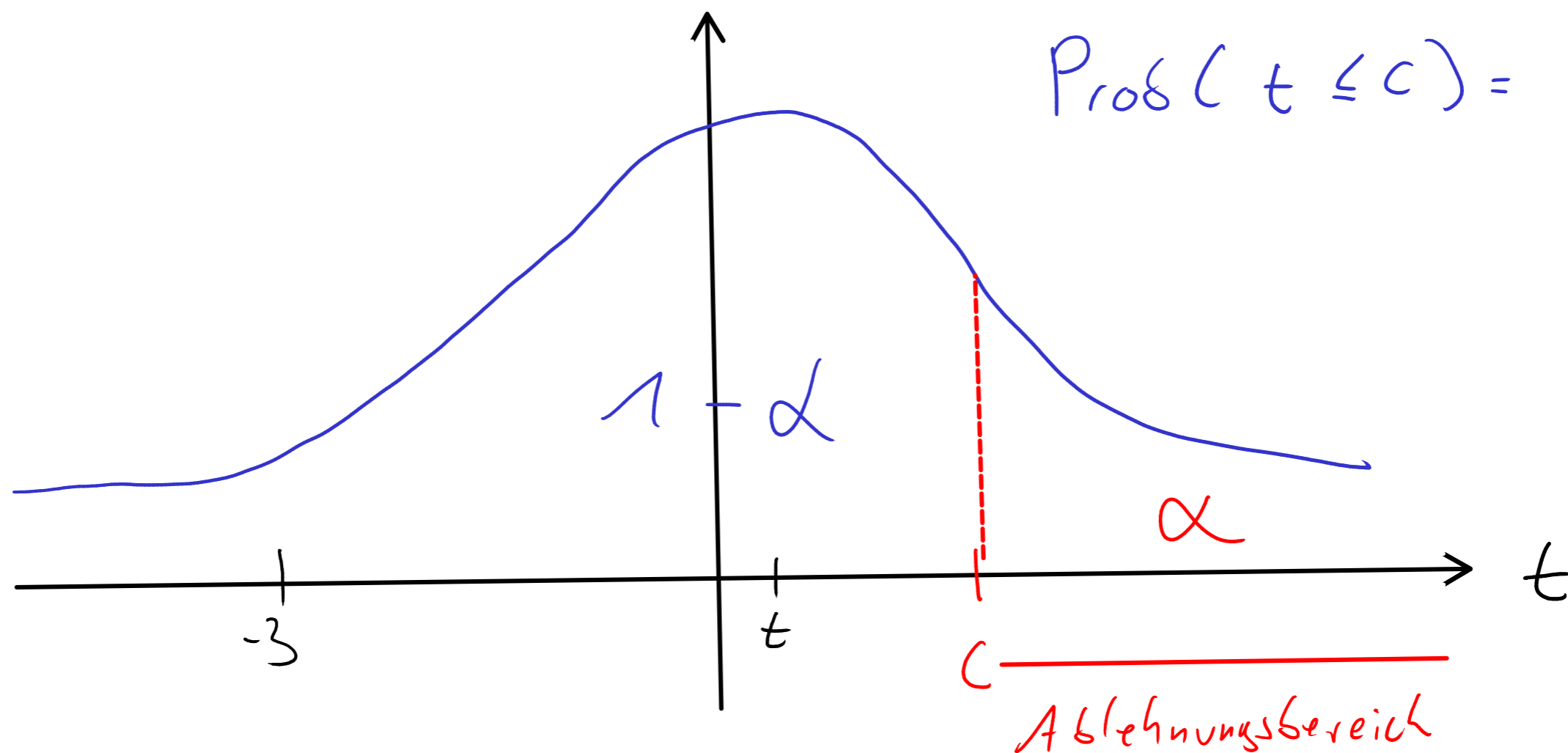
$$\tilde{H}_0: \beta_3 \leq 0$$

$$\tilde{H}_1: \beta_3 > 0$$

$$C_{t_{120}; 99\%} = 2,358$$

$$t_4 = \frac{\hat{\beta}_3 - 0}{se(\hat{\beta}_3)} = 0,44$$

$t_4 < C \rightarrow H_0$ nicht
verwerfen



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Jahr:	2007	2008	2009	2010	2011
Zeit:	10,03s	9,69s	9,58s	9,82s	9,76s
Jahr:	2012	2013	2014	2015	2016
Zeit:	9,63s	9,77s	9,98s	9,79s	9,81s

$$y_i = \beta_0 + u_i$$

$$\hat{\beta}_0 = \bar{y} \quad X = L$$

$$\text{Var}(\hat{\beta}_0) = \sigma^2 \cdot (L' L)^{-1} = \frac{1}{n} \sigma^2$$

$$= (L' L)^{-1} L' y = (n)^{-1} \sum y_i = \bar{y}$$

$$\widehat{\text{Var}}(\hat{\beta}_0) = \frac{1}{n} \hat{\sigma}^2$$

$$\hat{u} = y - \hat{y} = y - \hat{\beta}_0 = y - \bar{y}$$

$$\hat{\sigma}^2 = \frac{1}{n-k-1} \hat{u}' \hat{u}$$

$$\hat{u}' \hat{u} = \sum_{i=1}^n (y_i - \bar{y})^2$$

\uparrow
 $k=0$

$$= \frac{1}{n-1} \sum_{i=1}^n (y_i - \bar{y})^2$$

$$t_0 = \frac{\bar{y} - \gamma}{\sqrt{\frac{1}{n} \frac{1}{n-1} \sum_{i=1}^n (y_i - \bar{y})^2}}$$

#5

$$\gamma = 9,75$$

Modell 1: KQ, benutze die Beobachtungen 1-10

Abhängige Variable: best

	Koeffizient	Std.-fehler	t-Quotient	p-Wert
const	9,78600	0,0442016	221,4	3,98e-18 ***

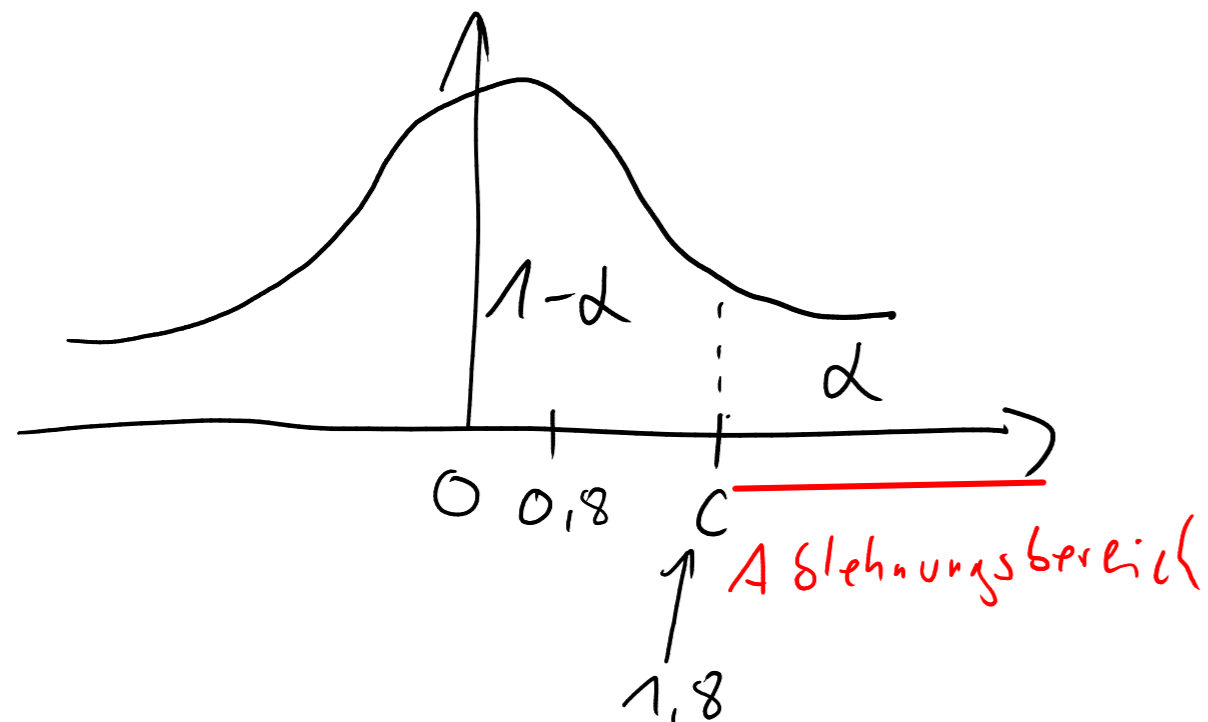
Mittel abhängige Var.	9,786000	Stdabw. abhängige Var.	0,139778
Summe quad. Residuen	0,175840	Stdfehler Regression	0,139778
R-Quadrat	0,000000	Korrigiertes R-Quadrat	0,000000
Log-Likelihood	6,014444	Akaike-Kriterium	-10,02889
Schwarz-Kriterium	-9,726303	Hannan-Quinn-Kriterium	-10,36082

$$H_0: \beta_0 \leq 9,75$$

$$H_1: \beta_0 > 9,75$$

aus Aufgabenstellung

$$t_0 = \frac{9,786 - 9,75}{0,0442016} = \frac{0,036}{0,0442016} = 0,81445$$



$$C_{t_{\beta_0}; 95\%} = 1,833$$

$$t_0 < C$$

$\Rightarrow H_0$ nicht ablehnen.